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ADAPTIVE TESTS *

by

K. S. Mak

Purdue University

Technical Report #88-38

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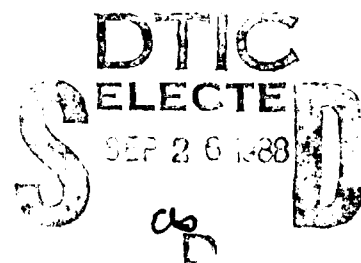
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August 1988

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Adaptive Tests

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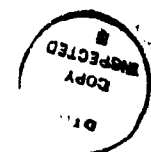
K. S. Mak¹

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Abstract

The problem of hypothesis testing when the distribution is specified only up to a nuisance parameter is considered. A test is said to be adaptive if it is asymptotically optimal regardless of the value of the nuisance parameter. The exponential rate of convergence to zero of the probability of type II error when the probability of type I error converges to zero exponentially fast at a fixed rate is used as the optimal criterion. A necessary and sufficient condition for the existence of adaptive test is obtained.

for $J_1(K_2)$



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1 Introduction

Let $\mathbf{x} = (x_1, \dots, x_n)$ be n independent, identically, distributed observations on a random variable X having distribution P or Q . It is desired to test the null hypothesis that X has distribution P versus the alternative that X has distribution Q .

Let $\phi_n = \phi_n(\mathbf{x})$ be any test function. Let A be a non-negative number. As in Tusnády (1977), a sequence of test functions $\{\phi_n\}$ is said to have exponential rate A if

$$\begin{aligned} \limsup_{n \rightarrow \infty} E_P \phi_n &< 1, \quad \text{when } A = 0, \\ \limsup_{n \rightarrow \infty} n^{-1} \log E_P \phi_n &\leq -A, \quad \text{when } A > 0. \end{aligned} \quad (1.1)$$

Let $\Phi_A(P)$ be the set of all sequences of tests that have exponential rate A . Let

$$B(A, P, Q) = -\inf \left\{ \liminf_{n \rightarrow \infty} n^{-1} \log E_Q(1 - \phi_n) : \{\phi_n\} \in \Phi_A(P) \right\}. \quad (1.2)$$

In other words, $B(A, P, Q)$ is the optimal exponential rate at which the probability of type II error can converge to zero. A sequence of tests $\{\phi_n\}$ in $\Phi_A(P)$ is said to be asymptotically optimal if

$$\liminf_{n \rightarrow \infty} n^{-1} \log E_Q(1 - \phi_n) = -B(A, P, Q).$$

Let the densities of P, Q be denoted by $p(x), q(x)$ respectively. It is assumed that $E_P(q(x)/p(x))^t < \infty$ for all t . It is known that (see Bahadur (1971)) if $\lim_{n \rightarrow \infty} E_P \phi_n = \alpha, 0 < \alpha < 1$ (so that the sequence of tests $\{\phi_n\}$ is of exponential rate 0), then $B(0, P, Q) = K(P, Q)$ where $K(P, Q) = E_P \log(p(x)/q(x))$ is the Kullback-Leibler information number.

When $A > 0$, it is shown in Blahut (1974), Tusnády (1977), and Birgé (1981) that

$$-B(A, P, Q) = \inf_{t > 0} \left\{ tC + \log E_Q \left(\frac{p(x)}{q(x)} \right)^t \right\} \quad (1.3)$$

where C is determined by

$$-A = \inf_{t > 0} \left\{ -tC + \log E_P \left(\frac{q(x)}{p(x)} \right)^t \right\}. \quad (1.4)$$

Futhermore, if $0 < A < K(Q, P)$, then $B = B(A, P, Q) > 0$ and

$$A = B + C. \quad (1.5)$$

A similar notion of asymptotic optimality for composite hypotheses has been investigated in Bahadur (1960), Hoeffding (1965), Brown (1971), Tusnády (1977) and Birgé (1981).

Now assume that the distributions P and Q are not determined exactly, but only up to a finite-valued nuisance parameter $\tau, \tau = 1, \dots, l$. For example, a message, in one of l possible languages using the same alphabets, is to be transmitted through a noisy channel n times and a choice has to be made between two possible messages or rather the probability distributions associated with the messages, without knowing which language is used. Another example is that there are l measurement types and for each type α , the measurement has two possible distributions P_α and Q_α . Thus, one has to test P_α versus Q_α with α unknown.

Let A_1, \dots, A_l be fixed positive numbers. A test ϕ_n^α is called an adaptive test if it is an asymptotically optimal test of rate A_α for each value of the nuisance parameter α . That is, ϕ_n^α is adaptive if for each α ,

$$\begin{aligned} \limsup_{n \rightarrow \infty} n^{-1} \log E_P^\alpha \phi_n^\alpha &\leq -A_\alpha \text{ and} \\ \liminf_{n \rightarrow \infty} n^{-1} \log E_Q^\alpha (1 - \phi_n^\alpha) &= -B(A_\alpha, P_\alpha, Q_\alpha). \end{aligned} \quad (1.6)$$

The existence of adaptive tests of rate 0 has been investigated in Rukhin (1982, 1986). A necessary and sufficient condition for the existence of such a test is that $K(P_\alpha, Q_\beta) \geq K(P_\beta, Q_\beta)$ for all α, β . It is shown in section 2 that for adaptive tests to exist, any two distributions $P_\alpha, Q_\beta, \alpha \neq \beta$ cannot be more 'difficult' to distinguish than P_α, Q_α . Here, 'difficulty' in distinguishing two distributions is measured by the rate of convergence of the type II error. When an adaptive test exists, it is shown that a weighted likelihood ratio test with weights depending on the rates of convergence of the type I and II errors is always adaptive. An overall maximum likelihood ratio test may not be adaptive.

2 Condition for the Existence of Adaptive Tests

Let $p_\alpha(x), q_\alpha(x)$ be the densities of P_α, Q_α respectively. Let A_1, \dots, A_l be positive constants and denote $B(A_\alpha, P_\alpha, Q_\beta), C(A_\alpha, P_\alpha, Q_\beta)$ by $B_{\alpha\beta}, C_{\alpha\beta}$ respectively. Also, let $\psi(t|P_\alpha, Q_\beta) = E_P^\alpha(q_\beta(x)/p_\alpha(x))^t$. Assume that $\psi(t|P_\alpha, Q_\beta) < \infty$ for all t , for all α, β . Note that $\psi(t|P_\alpha, Q_\beta)$ is the moment generating function of $\log(q_\beta(x)/p_\alpha(x))$, thus by the finiteness assumption, $\psi(t|P_\alpha, Q_\beta)$ is strictly convex

and differentiable, indeed differentiation can be carried out under the expectation sign. Let

$$T_n = T_n(\mathbf{x}) = n^{-1} \log \frac{\max_{\alpha} e^{nB_{\alpha\alpha}} \prod_{j=1}^n q_{\alpha}(x_j)}{\max_{\alpha} e^{nA_{\alpha}} \prod_{j=1}^n p_{\alpha}(x_j)}. \quad (2.1)$$

Let $\phi_n = \phi_n(\mathbf{x})$ be a test with critical region given by $T_n \geq 0$, i.e.

$$\phi_n = \begin{cases} 1 & \text{if } T_n \geq 0 \\ 0 & \text{otherwise.} \end{cases} \quad (2.2)$$

Theorem 2.1 Assume that $0 < A_{\alpha} < \min_{\beta} K(Q_{\beta}, P_{\alpha})$, for all α . If $B_{\beta\alpha} \geq B_{\alpha\alpha}$ for all β, α , then ϕ_n is an adaptive test.

Proof. We first show that ϕ_n has exponential rate A_{α} for every α . Since,

$$\begin{aligned} E_P^{\alpha} \phi_n &\leq P_{\alpha} \left(n^{-1} \log \frac{\max_{\beta} e^{nB_{\beta\beta}} \prod_{j=1}^n q_{\beta}(x_j)}{e^{nA_{\alpha}} \prod_{j=1}^n p_{\alpha}(x_j)} \geq 0 \right) \\ &\leq \sum_{\beta=1}^l P_{\alpha} \left(n^{-1} \log \frac{e^{nB_{\beta\beta}} \prod_{j=1}^n q_{\beta}(x_j)}{e^{nA_{\alpha}} \prod_{j=1}^n p_{\alpha}(x_j)} \geq 0 \right) \\ &\leq l \cdot \max_{\beta} P_{\alpha} \left(n^{-1} \sum_{j=1}^n \log \frac{q_{\beta}(x_j)}{p_{\alpha}(x_j)} \geq A_{\alpha} - B_{\beta\beta} \right), \end{aligned}$$

and

$$P_{\alpha} \left(n^{-1} \sum_{j=1}^n \log \frac{q_{\beta}(x_j)}{p_{\alpha}(x_j)} \geq A_{\alpha} - B_{\beta\beta} \right) \leq e^{-nt(A_{\alpha} - B_{\beta\beta})} \left[E_P^{\alpha} \left(\frac{q_{\beta}(x)}{p_{\alpha}(x)} \right)^t \right]^n$$

for any $t > 0$, it follows that,

$$E_P^{\alpha} \phi_n \leq l \cdot \max_{\beta} \inf_{t>0} e^{-nt(A_{\alpha} - B_{\beta\beta})} (\psi(t|P_{\alpha}, Q_{\beta}))^n.$$

Therefore,

$$\begin{aligned} \limsup_{n \rightarrow \infty} n^{-1} \log E_P^{\alpha} \phi_n &\leq \max_{\beta} \inf_{t>0} \{-t(A_{\alpha} - B_{\beta\beta}) + \log \psi(t|P_{\alpha}, Q_{\beta})\} \\ &\leq \max_{\beta} \inf_{t>0} \{-t(A_{\alpha} - B_{\alpha\beta}) + \log \psi(t|P_{\alpha}, Q_{\beta})\} \\ &= -A_{\alpha} \end{aligned}$$

as $A_{\alpha} - B_{\alpha\beta} = C_{\alpha\beta}$ by (1.5).

Since ϕ_n is a test of exponential rate A_α for each α , to show that ϕ_n is an adaptive test, it is enough to show that $\liminf_{n \rightarrow \infty} n^{-1} \log E_Q^\alpha(1 - \phi_n) \leq -B_{\alpha\alpha}$. Now,

$$E_Q^\alpha(1 - \phi_n) \leq l \cdot \max_\beta Q_\alpha \left(n^{-1} \sum_{j=1}^n \log \frac{p_\beta(x_j)}{q_\alpha(x_j)} > B_{\alpha\alpha} - A_\beta \right),$$

and so, by a similar argument, one obtains

$$\liminf_{n \rightarrow \infty} n^{-1} \log E_Q^\alpha(1 - \phi_n) \leq \max_\beta \inf_{t > 0} \{-t(B_{\alpha\alpha} - A_\beta) + \log \psi(t|Q_\alpha, P_\beta)\}. \quad (2.3)$$

Since the function $\log \psi(t|Q_\alpha, P_\beta)$ is strictly convex, there exist $0 < t_0 < 1$ such that $-B_{\beta\alpha} = t_0 C_{\beta\alpha} + \log \psi(t_0|Q_\alpha, P_\beta)$. Therefore,

$$\begin{aligned} & \inf_{t > 0} \{-t(B_{\alpha\alpha} - A_\beta) + \log \psi(t|Q_\alpha, P_\beta)\} \\ & \leq -t_0(B_{\alpha\alpha} - A_\beta) + \log \psi(t_0|Q_\alpha, P_\beta) \\ & = -B_{\alpha\alpha} + (1 - t_0)B_{\alpha\alpha} + t_0 A_\beta - B_{\beta\alpha} - t_0(A_\beta - B_{\beta\alpha}) \\ & = -B_{\alpha\alpha} + (1 - t_0)(B_{\alpha\alpha} - B_{\beta\alpha}) \\ & \leq -B_{\alpha\alpha}. \end{aligned}$$

Thus, from (2.3), $\liminf_{n \rightarrow \infty} n^{-1} \log E_Q^\alpha(1 - \phi_n) \leq -B_{\alpha\alpha}$, for all α and ϕ_n is an adaptive test. \square

A necessary condition for the existence of adaptive tests is obtained by considering the asymptotic behaviour of the most powerful test of the simple hypothesis $\prod_{j=1}^n p_\alpha(x_j)$ versus the simple alternative $w_\alpha \prod_{j=1}^n q_\alpha(x_j) + w_\beta \prod_{j=1}^n q_\beta(x_j)$, $\alpha \neq \beta$, where $w_k = e^{nb_k} / (e^{nb_\alpha} + e^{nb_\beta})$, $k = \alpha, \beta$, and b_α, b_β are real numbers.

For the remainder of this section, we assume that P_α, Q_α are absolutely continuous with respect to the Lebesgue measure for all α . Let \tilde{Q}_n denote the distribution with density $w_\alpha \prod_{j=1}^n q_\alpha(x_j) + w_\beta \prod_{j=1}^n q_\beta(x_j)$. Let ϕ_n^α be the following likelihood ratio test :

$$\phi_n^\alpha = \begin{cases} 1 & \text{if } \frac{w_\alpha \prod_{j=1}^n q_\alpha(x_j) + w_\beta \prod_{j=1}^n q_\beta(x_j)}{\prod_{j=1}^n p_\alpha(x_j)} \geq c_n \\ 0 & \text{otherwise,} \end{cases}$$

where c_n is a constant.

Lemma 2.1 For a fixed α , if $\lim_{n \rightarrow \infty} n^{-1} \log E_P^\alpha \phi_n^\alpha = -A_\alpha$, then

$$\lim_{n \rightarrow \infty} n^{-1} \log E_{\tilde{Q}_n}(1 - \phi_n^\alpha) = \max_{k=\alpha, \beta} (b_k + \rho_{\alpha k}(b_\alpha, b_\beta, C)) - \bar{b},$$

where $\rho_{\alpha k}(b_\alpha, b_\beta, C) = \inf_{s,t>0} \{s(C - b_\alpha) + t(C - b_\beta) + \log E_Q^k(\frac{p_\alpha(x)}{q_\alpha(x)})^s (\frac{p_\beta(x)}{q_\beta(x)})^t\}$ and $C = \max_{k=\alpha,\beta} (C_{\alpha k} + b_k)$, $\bar{b} = \max(b_\alpha, b_\beta)$.

Proof. We first show that the assumption implies that $\lim_{n \rightarrow \infty} c_n = C - \bar{b}$. From the definition of ϕ_n^α , one obtains

$$\begin{aligned} & \lim_{n \rightarrow \infty} n^{-1} \log E_P^\alpha \phi_n^\alpha \\ & \leq \liminf_{n \rightarrow \infty} n^{-1} \log P_\alpha \left(n^{-1} \log 2 \cdot \max_{k=\alpha,\beta} w_k \prod_{j=1}^n \frac{q_k(x_j)}{p_\alpha(x_j)} \geq c_n \right) \\ & \leq \max_{k=\alpha,\beta} \liminf_{n \rightarrow \infty} n^{-1} \log P_\alpha \left(n^{-1} \sum_{j=1}^n \log \frac{q_k(x_j)}{p_\alpha(x_j)} \geq c_n - b_k + n^{-1} \log \frac{w}{2} \right) \\ & \leq \max_{k=\alpha,\beta} \inf_{t>0} \{-t(\bar{C} - b_k + \bar{b}) + \log \psi(t|P_\alpha, Q_k)\} \end{aligned}$$

where $w = e^{nb_\alpha} + e^{nb_\beta}$, $\bar{C} = \limsup_{n \rightarrow \infty} c_n$. Therefore, $\inf_{t>0} \{-t(\bar{C} - b_k + \bar{b}) + \log \psi(t|P_\alpha, Q_k)\} \geq \inf_{t>0} \{-tC_{\alpha k} + \log \psi(t|P_\alpha, Q_k)\}$ for $k = \alpha$ or β . This implies that $\bar{C} - b_k + \bar{b} \leq C_{\alpha k}$, for $k = \alpha$ or β , i.e.

$$\bar{C} \leq \max_{k=\alpha,\beta} (C_{\alpha k} + b_k) - \bar{b}. \quad (2.4)$$

Also, $E_P^\alpha \phi_n^\alpha \geq P_\alpha(n^{-1} \log w_k \prod_{j=1}^n (q_k(x_j)/p_\alpha(x_j)) \geq c_n)$ for $k = \alpha$ and β . Thus

$$E_P^\alpha \phi_n^\alpha \geq \max_{k=\alpha,\beta} P_\alpha \left(n^{-1} \sum_{j=1}^n \log \frac{q_k(x_j)}{p_\alpha(x_j)} \geq c_n - b_k + n^{-1} \log w \right).$$

Using a similar argument as above, one obtains, for $k = \alpha$ and β ,

$$\inf_{t>0} \{-tC_{\alpha k} + \log \psi(t|P_\alpha, Q_k)\} \geq \inf_{t>0} \{-t(\underline{C} - b_k + \bar{b}) + \log \psi(t|P_\alpha, Q_k)\},$$

where $\underline{C} = \liminf_{n \rightarrow \infty} c_n$. Hence, $C_{\alpha k} \leq \underline{C} - b_k + \bar{b}$ for each k , i.e

$$\underline{C} \geq \max_{k=\alpha,\beta} (C_{\alpha k} + b_k) - \bar{b}. \quad (2.5)$$

Equations (2.4) and (2.5) imply that $\lim_{n \rightarrow \infty} c_n = C - \bar{b}$. Now,

$$\begin{aligned} E_Q^k(1 - \phi_n^\alpha) & \leq Q_k \left(n^{-1} \log \max_{k=\alpha,\beta} w_k \prod_{j=1}^n \frac{q_k(x_j)}{p_\alpha(x_j)} < c_n \right) \\ & = Q_k \left(n^{-1} \sum_{j=1}^n \log \frac{q_k(x_j)}{p_\alpha(x_j)} < c_n - b_k + n^{-1} \log w, k = \alpha, \beta \right) \end{aligned} \quad (2.6)$$

Similarly,

$$E_Q^k(1 - \phi_n^\alpha) \geq Q_k \left(n^{-1} \sum_{j=1}^n \log \frac{q_k(x_j)}{p_\alpha(x_j)} < c_n - b_k + n^{-1} \log \frac{w}{2}, k = \alpha, \beta \right). \quad (2.7)$$

Applying Theorem 5.1 of Groeneboom et al. (1979), (2.6) and (2.7) imply that

$$\lim_{n \rightarrow \infty} n^{-1} \log E_Q^k(1 - \phi_n^\alpha) = \rho_{\alpha k}(b_\alpha, b_\beta, C)$$

for $k = \alpha, \beta$. Since $E_{\tilde{Q}_n}(1 - \phi_n^\alpha) = w_\alpha E_Q^\alpha(1 - \phi_n^\alpha) + w_\beta E_Q^\beta(1 - \phi_n^\alpha)$, we have

$$\lim_{n \rightarrow \infty} n^{-1} \log E_{\tilde{Q}_n}(1 - \phi_n^\alpha) = \max_{k=\alpha, \beta} \lim_{n \rightarrow \infty} n^{-1} \log w_k E_Q^k(1 - \phi_n^\alpha)$$

and the result follows. \square

Since ϕ_n^α is the most powerful test of $\prod_{j=1}^n p_\alpha(x_j)$ versus $w_\alpha \prod_{j=1}^n q_\alpha(x_j) + w_\beta \prod_{j=1}^n q_\beta(x_j)$, this yields the following

Corollary 2.1 For any test ϕ_n^* such that $\limsup_{n \rightarrow \infty} n^{-1} \log E_P^\alpha \phi_n^* \leq -A_\alpha$, for a fixed α

$$\max_{k=\alpha, \beta} (b_k + \liminf_{n \rightarrow \infty} n^{-1} \log E_Q^k(1 - \phi_n^*)) \geq \max_{k=\alpha, \beta} (b_k + \rho_{\alpha k}(b_\alpha, b_\beta, C))$$

for any $\beta = 1, \dots, l$ and any b_1, \dots, b_l .

Corollary 2.2 $\rho_{\alpha k}(b_\alpha, b_\beta, C) \geq -B(A_\alpha, P_\alpha, Q_k), k = \alpha, \beta$.

Proof. As shown in the proof of lemma 2.1, $\lim_{n \rightarrow \infty} n^{-1} \log E_P^\alpha \phi_n^\alpha = -A_\alpha$ implies that $\lim_{n \rightarrow \infty} n^{-1} \log E_Q^k(1 - \phi_n^\alpha) = \rho_{\alpha k}(b_\alpha, b_\beta, C)$ for $k = \alpha, \beta$. Since, as a test of P_α versus Q_k , ϕ_n^α has exponential rate A_α , the result follows from (1.2). \square

Theorem 2.2 If an adaptive test exists, then for all α, β

$$B_{\alpha\alpha} \leq B_{\beta\alpha}.$$

Proof. Let ϕ_n^a be an adaptive test. Since ϕ_n^a satisfies the condition in corollary 2.1, by letting $b_\alpha = B_{\alpha\alpha}$ and $b_\beta = B_{\beta\beta}$, one obtains

$$0 = \max_{k=\alpha, \beta} (B_{kk} - B_{kk}) \geq \max_{k=\alpha, \beta} (B_{kk} + \rho_{\alpha k}(B_{\alpha\alpha}, B_{\beta\beta}, C)),$$

i.e. $\rho_{\alpha k}(B_{\alpha\alpha}, B_{\beta\beta}, C) \leq -B_{kk}$ for $k = \alpha, \beta$. But, from corollary 2.2, with $k = \beta$, $\rho_{\alpha\beta}(B_{\alpha\alpha}, B_{\beta\beta}, C) \geq -B_{\alpha\beta}$, so that $B_{\beta\beta} \leq B_{\alpha\beta}$. Since α, β are arbitrary, the result follows. \square

Hence from Theorems 2.1 and 2.2 we have established

Corollary 2.3 If P_α, Q_α are absolutely continuous with respect to the Lebesgue measure for all α , then an adaptive test exists iff for all α, β ,

$$B_{\alpha\alpha} \leq B_{\beta\alpha}. \quad (2.8)$$

From the proof of theorem 2, we see that if an adaptive test exists, then $\rho_{\alpha\beta}(B_{\alpha\alpha}, B_{\beta\beta}, C) \leq -B_{\beta\beta}$ for all α, β . Suppose that $P_\alpha = Q_\beta$ for some $\alpha \neq \beta$. Assume that A_β satisfies the condition in theorem 2.1, in particular $A_\beta < K(Q_\beta, P_\beta)$, i.e. $B_{\beta\beta} > 0$. Consider

$$\begin{aligned} & \rho_{\alpha\beta}(B_{\alpha\alpha}, B_{\beta\beta}, C) \\ &= \inf_{s, t > 0} \left\{ s(C - B_{\alpha\alpha}) + t(C - B_{\beta\beta}) + \log E_Q^\beta \left(\frac{p_\alpha(x)}{q_\alpha(x)} \right)^s \left(\frac{p_\alpha(x)}{q_\beta(x)} \right)^t \right\} \\ &\geq \inf_{s > 0} \left\{ s(C - B_{\alpha\alpha}) + \log E_Q^\beta \left(\frac{p_\alpha(x)}{q_\alpha(x)} \right)^s \right\} + \inf_{t > 0} \{ t(C - B_{\beta\beta}) \}. \end{aligned} \quad (2.9)$$

The second term of (2.9) is zero because $C - B_{\beta\beta} = \max_{k=\alpha, \beta} (C_{\alpha k} + B_{kk}) - B_{\beta\beta} \geq C_{\alpha\beta} > 0$ when $P_\alpha = Q_\beta$. Let $f(s)$ be the function in the paranthesis of the first term in (2.9). Then $f'(0) = C - B_{\alpha\alpha} + E_Q^\beta \log p_\alpha(x)/q_\alpha(x) = C - B_{\alpha\alpha} + K(P_\alpha, Q_\alpha) \geq C_{\alpha\alpha} + K(P_\alpha, Q_\alpha) \geq 0$. Since $f(s)$ is a convex function, this implies that $\inf_{s > 0} f(s) = 0$. That is $\rho_{\alpha\beta}(B_{\alpha\alpha}, B_{\beta\beta}, C) = 0$. But, by assumption $B_{\beta\beta} > 0$, and inequality $\rho_{\alpha\beta}(B_{\alpha\alpha}, B_{\beta\beta}, C) \leq -B_{\beta\beta}$ is impossible, and hence adaptive test cannot exist. This yields

Corollary 2.4 If $P_\alpha = Q_\beta$ for some $\alpha \neq \beta$, then an adaptive test does not exist.

Remarks. 1. If $P_\alpha = P$ for some P for all α , then (2.8) always holds, i.e. an adaptive test always exists.

2. By interchanging the roles of P_α, Q_α , we can define a similar notion of adaptation, i.e. a test is adaptive if the type II error converges to zero at a guaranteed rate while the type I error converges to zero at the optimal rate. Thus a test with the following properties:

$$\limsup_{n \rightarrow \infty} n^{-1} \log E_Q^\alpha(1 - \phi_n) \leq -A_\alpha \text{ and}$$

$$\liminf_{n \rightarrow \infty} n^{-1} \log E_P^\alpha \phi_n = -B(A_\alpha, Q_\alpha, P_\alpha)$$

exists iff

$$B(A_\alpha, Q_\alpha, P_\alpha) \leq B(A_\beta, Q_\beta, P_\alpha), \text{ for } \alpha, \beta = 1, \dots, l. \quad (2.10)$$

It is easy to see that $\lim_{A_\beta \rightarrow 0+} B(A_\beta, Q_\beta, P_\alpha) = K(Q_\beta, P_\alpha)$. Thus, by setting $A_1 = \dots = A_l = 0$, i.e. by letting the type II error converges to a positive constant, condition (2.10) becomes $K(Q_\alpha, P_\alpha) \leq K(Q_\beta, P_\alpha)$. This result is obtained in Rukhin (1986).

3. Corollary 2.3 can be extended to the case when the nuisance parameter has countably many values. Assume that $\inf_\beta K(Q_\beta, P_\alpha) > 0$ and $0 < A_\alpha < \inf_\beta K(Q_\beta, P_\alpha)$ for every α . Let $\{k_n\}$ be a non-decreasing sequence such that $n^{-1} \log k_n \rightarrow 0$ as $n \rightarrow \infty$. Consider the following test:

$$\phi_n^* = \begin{cases} 1 & \text{if } n^{-1} \log \frac{\max_{\beta \leq k_n} e^{nB_{\beta\beta}} \prod_{j=1}^n q_\beta(x_j)}{\max_{\beta \leq k_n} e^{nA_\beta} \prod_{j=1}^n p_\beta(x_j)} \geq 0 \\ 0 & \text{otherwise.} \end{cases}$$

Using corollaries 2.1 and 2.2, it is clear that if an adaptive test exists, then (2.8) holds for all α, β . We claim that ϕ_n^* is an adaptive test if (2.8) holds. We first show that ϕ_n^* is a test of rate A_α for every α . Pick n large enough so that $\alpha \leq k_n$, then

$$E_P^\alpha \phi_n^* \leq k_n \cdot \max_{\beta \leq k_n} \inf_{t > 0} e^{-nt(A_\alpha - B_{\beta\beta})} (\psi(t|P_\alpha, Q_\beta))^n.$$

From the proof of theorem 2.1. $\inf_{t > 0} \{e^{-t(A_\alpha - B_{\beta\beta})} \psi(t|P_\alpha, Q_\beta)\} \leq e^{-A_\alpha}$, for any β . Thus,

$$\begin{aligned} \limsup_{n \rightarrow \infty} n^{-1} \log E_P^\alpha \phi_n^* &\leq \lim_{n \rightarrow \infty} n^{-1} \log k_n - A_\alpha \\ &= -A_\alpha. \end{aligned}$$

Now, we show that type II error of the test ϕ_n^* converges to zero at the optimal rate. Using the same argument as above

$$\begin{aligned} &\liminf_{n \rightarrow \infty} n^{-1} \log E_Q^\alpha (1 - \phi_n^*) \\ &\leq \sup_\beta \inf_{t > 0} \{-t(B_{\alpha\alpha} - A_\beta) + \log \psi(t|Q_\alpha, P_\beta)\} + \lim_{n \rightarrow \infty} n^{-1} \log k_n \\ &\leq -B_{\alpha\alpha}. \end{aligned}$$

Hence, ϕ_n^* is an adaptive test.

4. When an adaptive test exists, then the test defined in (2.2) is always adaptive. However, an overall maximum likelihood ratio test, i.e. a test with critical region $\{x : \max_\alpha \prod_{j=1}^n q_\alpha(x_j) / \max_\alpha \prod_{j=1}^n p_\alpha(x_j) \geq e^{nc_n}\}$, for some constant c_n , is not necessarily adaptive even when an adaptive test exists. Let $\hat{\phi}_n$ denote such a

test. If $c_n \equiv \max_{k=\alpha,\beta} C_{\alpha\beta}$, then it can be shown that $\hat{\phi}_n$ is a test of rate A_α for each α . If $A_\alpha, \alpha = 1, \dots, l$ are picked such that $C_{11} = \dots = C_{ll}$, then a sufficient condition for the test $\hat{\phi}_n$ to be adaptive is that $B_{\alpha\beta} \geq B_{\beta\beta} + t_\alpha(C - C_{\alpha\beta})$ where $C = \max_{\alpha,\beta} C_{\alpha\beta}$, $t_\alpha = \max_\beta t_{\beta\alpha}$, and $t_{\beta\alpha}$ is the point where $-tC_{\alpha\beta} + \log \psi(t|P_\alpha, Q_\beta)$ attains its minimum.

3 Example

Let P_α, Q_α belong to an exponential family with densities p_α, q_α given by

$$p_\alpha(x) = \exp\{\xi'_\alpha x - \chi(\xi_\alpha)\},$$

$$q_\alpha(x) = \exp\{\eta'_\alpha x - \chi(\eta_\alpha)\}.$$

Let $g(t, u) = -tu + \log \psi(t|P_\alpha, Q_\beta)$. By a straightforward calculation,

$$\begin{aligned} \psi(t|P_\alpha, Q_\beta) &= \int \exp t\{\eta'_\beta x - \xi'_\alpha x - \chi(\eta_\beta) + \chi(\xi_\alpha)\} \exp\{\xi'_\alpha x - \chi(\xi_\alpha)\} dx \\ &= \exp\{-(1-t)\chi(\xi_\alpha) - t\chi(\eta_\beta) + \chi(t\eta_\beta + (1-t)\xi_\alpha)\}. \end{aligned}$$

Thus, $g(t, u) = -tu - (1-t)\chi(\xi_\alpha) - t\chi(\eta_\beta) + \chi(t\eta_\beta + (1-t)\xi_\alpha)$. Therefore we have

$$g'(t) = -u + \chi(\xi_\alpha) - \chi(\eta_\beta) + \chi'(t\eta_\beta + (1-t)\xi_\alpha),$$

where g', χ' are derivatives of g, χ with respect to t .

If $s_{\alpha\beta}$ satisfies

$$s_{\alpha\beta}\chi'(s_{\alpha\beta}\eta_\beta + (1-s_{\alpha\beta})\xi_\alpha) - \chi(s_{\alpha\beta}\eta_\beta + (1-s_{\alpha\beta})\xi_\alpha) + \chi(\xi_\alpha) = A_\alpha, \quad (3.1)$$

and let

$$C_{\alpha\beta}(A_\alpha) = \chi(\eta_\beta) - \chi(\xi_\alpha) - \chi'(s_{\alpha\beta}\eta_\beta + (1-s_{\alpha\beta})\xi_\alpha), \quad (3.2)$$

then $g(s_{\alpha\beta}, C_{\alpha\beta}) = -A_\alpha$ and $g'(s_{\alpha\beta}, C_{\alpha\beta}) = 0$.

Suppose $0 < A_\alpha < \min_\beta K(Q_\beta, P_\alpha)$, for each α , then an adaptive test exists iff

$$\begin{aligned} A_\alpha - \chi(\eta_\beta) + \chi(\xi_\alpha) + \chi'(s_{\alpha\beta}\eta_\beta + (1-s_{\alpha\beta})\xi_\alpha) \\ \geq A_\beta - \chi(\eta_\beta) + \chi(\xi_\beta) + \chi'(s_{\beta\beta}\eta_\beta + (1-s_{\beta\beta})\xi_\beta) \end{aligned}$$

or,

$$\chi(\xi_\alpha) - \chi(\xi_\beta) + \chi'(s_{\alpha\beta}\eta_\beta + (1-s_{\alpha\beta})\xi_\alpha) - \chi'(s_{\beta\beta}\eta_\beta + (1-s_{\beta\beta})\xi_\beta) \geq A_\beta - A_\alpha. \quad (3.3)$$

Suppose $P_\alpha \sim N(\theta_\alpha, 1)$, $Q_\alpha \sim N(\mu_\alpha, 1)$, then $K(Q_\beta, P_\alpha) = 2^{-1}(\mu_\beta - \theta_\alpha)^2$ and

$$\begin{aligned}\chi(t\eta_\beta + (1-t)\xi_\alpha) &= \frac{1}{2}(t\mu_\beta + (1-t)\theta_\alpha)^2, \\ \chi'(t\eta_\beta + (1-t)\xi_\alpha) &= (t\mu_\beta + (1-t)\theta_\alpha)(\mu_\beta - \theta_\alpha).\end{aligned}$$

It is easy to check that $s_{\alpha\beta} = \sqrt{\frac{2A_\alpha}{(\mu_\beta - \theta_\alpha)^2}}$ satisfies (3.1), and it follows from (3.2) that

$$C_{\alpha\beta}(A_\alpha) = \frac{(\mu_\beta - \theta_\alpha)^2}{2} \left(\frac{2\sqrt{2A_\alpha}}{|\mu_\beta - \theta_\alpha|} - 1 \right).$$

Thus, if $0 < A_\alpha < \min_\beta 2^{-1}(\mu_\beta - \theta_\alpha)^2$, then an adaptive test exists iff for all $\alpha \neq \beta$

$$(\mu_\alpha - \theta_\alpha)^2 \left(\frac{\sqrt{2A_\alpha}}{|\mu_\alpha - \theta_\alpha|} - \frac{1}{2} \right) - (\mu_\alpha - \theta_\beta)^2 \left(\frac{\sqrt{2A_\beta}}{|\mu_\alpha - \theta_\beta|} - \frac{1}{2} \right) \geq A_\beta - A_\alpha. \quad (3.4)$$

For example, when $\theta_1 = -1, \mu_1 = 1; \theta_2 = -1.5, \mu_2 = 0.5, A_1 = 0.5, A_2 = 1.125$, then $C_{11} = 0, C_{12} = 0.375, C_{21} = 0.625, C_{22} = 1$ and the above condition is satisfied. (Here $B_{11} = B_{21} = 0.5, B_{22} = B_{12} = 0.125$).

Consider the following overall maximum likelihood ratio test with the above parametric values:

$$\hat{\phi}_n(\mathbf{x}) = \begin{cases} 1 & \text{if } n^{-1} \log \frac{\max_{\alpha=1,2} \prod_{j=1}^n q_\alpha(z_j)}{\max_{\alpha=1,2} \prod_{j=1}^n p_\alpha(z_j)} \geq C \\ 0 & \text{otherwise.} \end{cases}$$

If the critical constant C is chosen such that $C = \max_{\alpha\beta} C_{\alpha\beta}(A_\alpha) = 1$, then $\hat{\phi}_n$ is a test of exponential rate A_α for $\alpha = 1, 2$. Now, consider the rate of convergence of the probability of type two error:

$$\lim_{n \rightarrow \infty} n^{-1} \log E_Q^\alpha(1 - \hat{\phi}_n) = \max_{k=1,2} \inf_{s,t>0} \left\{ (s+t) + \log E_Q^\alpha \left(\frac{p_k(x)}{q_1(x)} \right)^s \left(\frac{p_k(x)}{q_2(x)} \right)^t \right\}. \quad (3.5)$$

When $\alpha = 2$,

$$\begin{aligned}E_Q^2 \left(\frac{p_k(x)}{q_1(x)} \right)^s \left(\frac{p_k(x)}{q_2(x)} \right)^t &= \exp -\frac{1}{2} \{ s(1-s)(\theta_k - \mu_1)^2 + t(1-t)(\theta_k - \mu_2)^2 + \\ &\quad 2t(\theta_k - \mu_1)(\mu_1 - \mu_2) - 2st(\theta_k - \mu_2)(\theta_k - \mu_1) \}.\end{aligned}$$

Consider $k = 1$ and let the expression in the parenthesis of (3.5) be $g(s, t)$, i.e.

$$g(s, t) = s + t - \frac{9}{8}t(1-t) - 2s(1-s) + s + 3st.$$

By differentiating g with respect to s and t , one can show that $(\partial/\partial s)g$ and $(\partial/\partial t)g$ do not vanish simultaneously and therefore

$$\inf_{s,t>0} g(s,t) = \min\{\inf_{s>0} g(s,0), \inf_{t>0} g(0,t)\}.$$

By a straightforward calculation, $\inf_{s,t>0} g(s,t) \approx -.0035$. Hence,

$$\lim_{n \rightarrow \infty} n^{-1} \log E_Q^2(1 - \hat{\phi}_n) > -.125 = -B_2,$$

and it follows that $\hat{\phi}_n$ is not an adaptive test.

Note that in this example, the critical constant $C = 1$ is the 'best' in the sense that $\hat{\phi}_n$ is a test of exponential rate A_α to reach α . If C is replaced by c_n in the definition of $\hat{\phi}_n$ such that $\liminf_{n \rightarrow \infty} c_n = C' < 1$, then by a similar calculation as above, one can show that $\limsup_{n \rightarrow \infty} n^{-1} \log E_P^2 \hat{\phi}_n > -A_2$.

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